Time-dependent cross-correlations between different stock returns: A directed network of influence

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We study the time-dependent cross-correlations of stock returns, i.e., we measure the correlation as the function of the time shift between pairs of stock return time series using tick-by-tick data. We find a weak but significant effect showing that in many cases the maximum correlation appears at nonzero time shift, indicating directions of influence between the companies. Due to the weakness of this effect and the shortness of the characteristic time (of the order of a few minutes), our findings are compatible with market efficiency. The interaction of companies defines a directed network of influence.

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I. INTRODUCTION

In the risk minimization of portfolio optimization, it is very important to consider how the returns of different companies correlate with each other. For this purpose the study of the equal time cross-correlations between stocks has attracted much interest [1,2]: The clustering properties and the comparison between the time and ensemble averages have provided much useful information in this respect.

In statistical physics *time-dependent* correlations are also of major interest. Due to their role in the fluctuation dissipation theorem they constitute as the main tool for determining transport coefficients. The famous Onsager reciprocity relations have their roots in the symmetry properties of the timedependent cross-correlations [3]. Obviously, in an economic system there is no reason to assume that the time reversal symmetry or detailed balance is maintained. Nevertheless, it is of interest to investigate the time-dependent crosscorrelations between stock returns because they contain information about the way the prices influence each other, which are the dominant stocks, and to what extent this dominance is reflected in the price changes under the conditions of an efficient market.

The time-dependent correlations between the indices of different stock exchanges were already studied empirically by Refs. [4,5] and with a microscopic model by Ref. [6]. They showed that there exists a time shift in the cross-correlations that arises from the fact that the different stock markets are open in different time cycles during the day as the Earth rotates.

In this paper we study time-dependent cross-correlation functions of the returns of different stocks taken from the New York Stock Exchange (NYSE). As we will show, in many cases the maximum of the correlation as a function of time is not at zero but shifted, meaning that there exists some "pulling" effect between the companies, i.e., one of them influences the price behavior of the others. However, this effect cannot be strong and the shift should be small, otherwise the effect could be utilized for arbitrage purposes, which is excluded from an efficient market. In fact, the investigated cases do not contradict with these criteria. The time-dependent correlation between the stocks was studied before by Refs. [7,8]. The results of Ref. [7] seems to contradict with ours because they studied weekly returns, and they found significant cross-correlations on the weekly scale. However, the results agree considering the "pulling" effect, namely, that the cross-correlations are asymmetric.

The paper is organized as follows: In the following section we give a short description of the dataset. In Sec. III we present the method of analysis and show how it works on an artificial set of data. Sec. IV is devoted to the presentation of the results. The paper terminates with a discussion.

II. DATA

One of the stylized facts of markets is that the autocorrelation of stock returns decays exponentially with a very short characteristic correlation time, which is in the range of a few minutes [9,10]. This is understood as a signature of market efficiency [11]. Since cross-correlations could also be used for arbitrage, one should not expect effects much beyond the above scale and therefore high frequency data are needed. We have analyzed the Trade and Quote (TAQ) database for N=54 days over the time period from 01.12.1997 to 09.03.1998, which includes tick-by-tick data for about 10 000 companies. Since this is quite a short time period we selected only those companies that were traded more than 15 000 times such that the number of companies reduced to 195.

Having these 195 time series we have to face the following problem: since the tradings do not happen simultaneously, the values of the returns have also to be defined for the time intervals between the tradings. According to the rules of the stock exchanges we have considered the price as constant between two changes. The whole trading time Tduring 1 day is divided into n small intervals or windows of size $\Delta t = T/n$. If the trading happens in the interval t the return takes the value

$$r_{\Delta t}(t) = \frac{\ln[p(t)]}{\ln[p(t-\Delta t)]},$$

where for simplicity the day index i is not indicated; otherwise it is zero.



FIG. 1. Illustration of how the correlation as the function of the time shift τ depends on the time deference Δt of the return. The correlation was measured on the artificial datasets with parameters $\rho = 0.01$, $\tau_0 = 200$, $\sigma = 1000$, $\alpha = 0.99$. The figures show that while for $\Delta t = 1$ no peak in the correlation function can be identified, by increasing the time difference to $\Delta t = 10$ the peak at $\tau = 200$ appears.

In order to avoid the problem of major return values stemming from the differences between opening and previous day's closing prices we simply took the days as independent, i.e., the averaging is separated into two steps: Over the intraday trading time T and over the trading days. The data prepared in this way were then analyzed from the point of view of time-dependent cross-correlations.

III. METHOD OF MEASURING THE CORRELATION

As mentioned in the Introduction we want to investigate the correlation of returns as the function of the time shift between pairs of stocks' return time series. The definition of the time-dependent correlation function $C_{A,B}(\tau)$ is

$$C_{\Delta t}^{A,B}(\tau) = \frac{\langle r_{\Delta t}^{A}(t) r_{\Delta t}^{B}(t+\tau) \rangle - \langle r_{\Delta t}^{A}(t) \rangle \langle r_{\Delta t}^{B}(t+\tau) \rangle}{\sigma_{A} \sigma_{B}}, \quad (1)$$

where $\sigma^2 = \langle [r_{\Delta t}(t) - \langle r_{\Delta t}(t) \rangle]^2 \rangle$ is the variance of the return. The notation $\langle \cdot \rangle$ means averaging over the whole trading time *T* and important details of this process will be given in the following.

Since the smallest interval between two tradings is one second, then $\Delta t = 1$ s seems to be a natural choice. However, for such a short window it quite often happens that at a given time step there is no transaction for one of the stocks (or for both) such that the return results in a zero contribution to the total correlations. Since the number of nonzero contributions is small, the correlation coefficients as a function of the time shift τ will strongly fluctuate. To avoid this problem one has to enlarge the time difference Δt and average the correlations over the starting points of the returns. In this way the average in Eq. (1) means the following:

$$\langle r_{\Delta t}^{A}(t) r_{\Delta t}^{B}(t+\tau) \rangle = \frac{1}{T} \sum_{t_{0}=0}^{\Delta t-1} \sum_{k=1}^{T/\Delta t} r_{\Delta t}^{A}(t_{0}+k\Delta t)$$
$$\times r_{\Delta t}^{B}(t_{0}+k\Delta t+\tau),$$
(2)

where the first sum runs over the starting points of the returns and the second one runs over the Δt wide windows of the returns.

In order to illustrate the effect that by taking larger time difference it is easier to identify the peaks in the correlation function—in other words, to locate the time that gives the maximal correlation—we simulated two series of artificial data sets. The first is a one-dimensional persistent random walk (RW) [12], which deviates from a normal RW by the fact that the probability α that it jumps in the same direction as in the previous step is higher than 0.5, i.e., the random walker remembers its history. The probability of an increment $x(t) \in \{\pm 1\}$ at time *t* is

$$P(x(t)) = \alpha \,\delta_{x(t), x(t-1)} + (1-\alpha)(1-\delta_{x, x(t-1)}). \tag{3}$$

The other time process is simply generated from the first one by shifting it by τ_0 and adding to it Gaussian random noise with zero mean and width σ :

$$y(t) = x(t - \tau_0) + \xi(t), \quad \xi \in N(0,\sigma).$$
 (4)

The advantage of this model is that the correlation function can be calculated analytically and the position of the maximum correlations can be adjusted at τ_0 ,

$$C(\tau) = \frac{(2\alpha - 1)^{|\tau - \tau_0|}}{\sqrt{\sigma + 1}}.$$
 (5)

After generating the two datasets we randomly drop points from both sets and keep only the fraction ρ of the points in order to have the same problem as with the original datasets that the jumps do not occur at the same time in the different time series. It is apparent from Fig. 1 that increasing the time difference, Δt helps identifying the time of maximum correlation. The fact that we dropped random points from the original data changes slightly the position of the maximum correlation as compared to Eq. (5).

Figure 1 shows that the decay of the correlation function is not exponential as in Eq. (5) but it decays approximately linearly down to the noise level. This is due to the averaging procedure we use with the increased time difference Δt . The correlation corresponding to larger time difference $C_{\Delta t}(\tau)$ can be written as the weighted sum of the one-step correlation functions C_1 that belongs to $\Delta t = 1$ s,

$$\langle r^{A}_{\Delta t}(t)r^{B}_{\Delta t}(t+\tau)\rangle = \left\langle \sum_{i=1}^{\Delta t} \delta^{A}(t+i)\sum_{i=1}^{\Delta t} \delta^{B}(t+i+\tau) \right\rangle$$
$$= C_{1}(\tau-\Delta t+1)+\dots+(\Delta t-1)$$
$$\times C_{1}(\tau-1)+\Delta t C_{1}(\tau)$$
$$+\dots+1 C_{1}(\tau+\Delta t-1), \qquad (6)$$

where $\delta = r_{\Delta t=1}$ is the return belonging to one second time difference.

Changing τ in Eq. (6) means changing the weights of the one-step correlation functions. Since the correlation function of the original datasets, see Eq. (5), decays exponentially, the maximum $C_1(\tau_0)$ will give the main contribution to the sum in Eq. (6) and because its weight is linear in τ , then $C_{\Delta t}(\tau)$ will decay approximately linearly. [It should be noted that the normalization factor in Eq. (1) does not change this consideration since it is independent of τ .]

There is only one question left, namely, how can we choose a smaller value for τ than for Δt ? The timedependent cross-correlation of the returns contains a product of the *return* of company A with that of company B shifted by τ . As the return is defined with the window Δt the values of τ could only be multiples of Δt . The solution is simply that one shifts the starting point of the return of company B by τ , as evident in Eq. (2), i.e., we make the time shift in the price function and in this way allow any time shift larger than the minimum trading time.

The above arguments of averaging give support to choose a value for Δt that is larger than the minimum trading time. However, it should not be too large since the averaging leads then to the smearing out of the maximum. As the width of the one-step correlation should be a few minutes, much larger time difference would mean that in the sum of Eq. (6) we mainly have terms, which are only due to noise. This suggests that the optimal choice for Δt is of the order of magnitude of one minute.

IV. RESULTS

As mentioned in Sec. II we have studied the correlation of 195 companies, which were traded during the available 54 days more than 15 000 times. In accordance with the arguments presented in the preceding section we have used $\Delta t = 100$ but checked that the results are quite robust within the range $50 \le \Delta t \le 500$. As already mentioned, we averaged over the starting points of the returns. For the maximum of the time shift we choose 2000 s. This is definitely beyond any reasonable characteristic time for correlations in return values because of market efficiency. In fact, using such a large value for the time shift allows us to measure the noise level, which the possible effect should be compared with.

For the resulting $195 \times 194/2$ correlation functions we measured the maximum value C_{max} the position τ_{max} at which time shift this maximum was found, and the ratio *R* of the maximum and strength of the noise defined as the variance of the correlation values for time shift values between 600 and 2000 s. We looked at those pairs of companies for



FIG. 2. Example for the measured shifted-time correlation function. The two companies are Ensco International (ESV) and Exxon Corp. (XON). The maximum correlation value is at -100 s, which means that the return time series of ESV has to be shifted back in order to get the maximal correlation, i.e., the price changes happen later in time. In other words, ESV is pulled by XON.

which these three values exceeded a prescribed threshold values, which we defined for $\Delta t = 100$ as $\tau_{max} \ge 100$, $C_{max} \ge 0.04$, $R \ge 6.0$. One example of the measured correlation function can be seen in Fig. 2. In this case the company XON (Exxon)—which is a large oil company—"pulls" the ESV (Ensco International) which provides drilling service to oil and gas companies. This effect is quite weak but the large value of *R* shows that it is significant. (See Table I for NYSE company abbreviations.)

The maximal value of the correlations turn out to be quite small, in average less than 0.1, (e.g., see Fig. 2), although the generally quoted equal time cross-correlations have much larger values. The root of this effect lies in the choice of the time difference, Δt . Increasing Δt increases the values of the equal time correlations [13].

In some cases the position of the maximum correlation was found at values much larger than few minutes that would be inconsistent with the efficient market behavior. A closer inspection revealed that in such cases the peak in the correlation function is caused by two major return values in the considered time series. The contribution of their product to the correlation-at appropriate value of the time shiftdominates the maximum of the correlation function. These are not the effects we are looking for, therefore we did not take them into account. In order to check whether the peak in the correlation is due to some single large return value or due to persistent influence of one of the stocks on the other we also studied how the correlation changes if the analyzed time window changes. We measured the shifted time correlation also for the first and for the second half of the given 3-month period and studied whether the correlation function remains qualitatively the same.

We also measured the correlation for shorter and for larger time difference, i.e., $\Delta t = 50$ and $\Delta t = 200$, respectively, because it may happen that by changing the time difference also the position of the maximal correlation value changes due to the averaging procedure described in Eq. (6). This can happen if the time-dependent correlation function for Δt has an asymmetric peak; see Fig. 3. Let us suppose that the left hand side is higher than the right hand one. For $\Delta t' > \Delta t$ the maximum will be shifted towards left as it can

TABLE I.	Company	names	and	description.
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Symbol	Name	Description
ARC	Atlantic Richfield Co.	Petroleum refining
AUD	Automatic Data Processing	Data communications and information services
BLS	Bellsouth Corp.	Telephone communication
BMY	Bristol-Myers Squibb Co.	Pharmaceutical preparations
CCI	Citicorp	Banking
CD	Cendant Corp.	Travel, real estate, vehicle, and financial services
CDG	Cliffs Drilling Co.	International drilling company
CHV	Chevron Corp.	Energy and chemical company, petroleum refining
CPQ	Compaq Computer Corp.	Electronic computers
DIS	Walt Disney Co.	Entertainment company
EK	Eastman Kodak Co.	Photography
ESV	Ensco International Inc.	Drilling oil and gas wells
EVI	Energy Ventures Inc.	Oil and gas field machinery
FCN	First Chicago NBD Corp.	Banking
FLC	Falcon Drilling Co. Inc.	Marine-based drilling
GE	General Electric Co.	Electronics, machinery
GLM	Global Marine Inc.	Drilling oil and gas wells
GLX	Glaxo Wellcome Plc.	Pharmaceutical preparations
GTW	Gateway 2000 Inc.	Electronic computers
HAL	Halliburton Co.	Oil field services
HD	Home Depot Inc.	Home improvement retailer
HWP	Hewlett-Packard Co.	Computers
IBM	International Business Machines Corp.	Computers
IP	International Paper Co.	Paper
JNJ	Johnson & Johnson	Health care products
JPM	Morgan J.P. Co. Inc.	Banking
KO	Coca-Cola Co.	Soft drinks
LEH	Lehman Brothers Holdings	Financial services
MOB	Mobil Corp.	Petroleum refining
MOT	Motorola Inc.	Semiconductor technology
MRK	Merck & Co Inc.	Pharmaceutical preparations
MU	Micron Technology Inc.	Semiconductor technology
NE	Noble Drilling Corp.	Drilling oil and gas wells
NN	Newbridge Networks Corp.	Telephone and telegraph apparatus
NOKA	Nokia Corp.	Mobile phones
PFE	Pfizer Inc.	Pharmaceutical preparations
PG	Procter & Gamble Co.	Soap and other detergents
RD	Royal Dutch Petroleum Comp.	Petroleum refining
SBH	Smithkline Plc	Pharmaceutical preparations
SLB	Schlumberger Limited LTD	Oil and gas field services
SUB	Summit Bank Corp.	Banking
TBR	Telecomunicacoes Brasileiras S.A.	Telecommunications
TER	Teradyne Inc.	Electrical instruments
TMX	Telefonos de Mexico	Telephone communication
TRV	Travelers Group Inc.	Fire, marine and casualty insurance
UAL	UAL Corp.	Air transportation
VRC	Varco International Inc.	Oil and gas field services
WDC	Western Digital Corp.	Computer storage devices
WLA	Warner Lambert Co.	Pharmaceutical preparations
WMI	wai-Mart Stores Inc.	Retail - variety stores
AUN	Exxon Corp.	Petroleum refining



FIG. 3. Example for a pair of companies for which the correlation function has an asymmetric peak. The curve with circles belongs to $\Delta t = 100$, the other with the squares to $\Delta t = 500$. The maximum of the second curve is at smaller time value because the left side of the peak—in the case of the curve with circles—is higher.

be shown through simple examples using Eq. (6). In the case of Fig. 2 the correlation function is also asymmetric but not at its peak (not near the maximum), which means that the maximum will not be shifted by increasing the time difference Δt .

The results show that the characteristic time shift is around 100 s, which is consistent with the effective market hypothesis. A time shift larger than the characteristic time of the decay of the return autocorrelations would contradict with the efficient market picture and could be used to arbitrage.

In general the more frequently traded companies are influencing ("pulling") the less frequently traded ones. This is not surprising since obviously the more frequently traded companies are more important. It is therefore more likely that they influence a smaller company than the other way around. Although this is the generic situation, there are a few exceptions when a less often traded company "pulls" the other one. In this study we found that in general one "small" company is influenced by many "large" companies and one "large" company pulls many "small" ones. As can be seen in Fig. 4 this behavior can be represented as a graph of directed links, where there are nodes from which many links go out (meaning that this node is influenced by many others) and there are other nodes where many links go in (these are the big companies influencing the less important ones).

V. DISCUSSION

In this paper we have analyzed the time-dependent crosscorrelation functions of the returns of stocks at the NYSE. We have studied whether there exists any pulling effect between stocks, i.e., whether at a given time the return value of one stock influences that of another stock at a different time.

In general we can see two types of mechanisms to generate significant correlation between two stocks.

(i) Some external effect (e.g., economic, political news, etc.) that influences both stock prices simultaneously. In this case the change for both prices appears at the same time, and the maximum of the correlation is at zero time shift.

(ii) One of the companies has an influence to the other (e.g., one of the company's operation depends on the other). In this case the price change of the influenced stock appears later in time because it needs some time to react on the price change of the first stock, in other words one of the stocks pulls the other. This pulling effect has been the main focus of our study in this article.

Since the correlation between stocks was expected to be small and the available set of data was somewhat limited we had to do a careful analysis. For this reason and test purposes we generated an artificial dataset with which we showed that by increasing the time window of the returns and by averaging over their starting points the detection of the correlation effect gets easier.



FIG. 4. Representation of the pulling effect between the companies. The direction of the arrows shows which company is pulling the other. The companies which appear in the figure show the most significant effects. With the real data we saw that it is possible to find pairs of stocks where the pulling effect exists, though it turned out to be small. In addition, the characteristic time shift—given by the position of maximum correlation—was found to be of the order of a few minutes. These findings are compatible with the efficient market picture.

As for the pulling effect we found that generically the more traded, and thus more important companies pull the relatively smaller companies. This result is consistent with that of [7]. In this light it is not surprising that in the study of the time-dependent cross-correlation functions of pairs of companies in the Dow Jones industrial average index no pulling effect was found. This underlines the fact that the Dow Jones companies are indeed among the most important stocks of the New York Stock Exchange.

Finally we would like to propose that although the ob-

served pulling effect was small, our careful analysis could show that it is significant for a considerable set of pairs of companies. We think that this property of the stock market should be added to the so called stylized facts. Of course, further analysis on more extensive data is needed to clarify further details of the time-dependent cross-correlations.

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- [1] R.N. Mantegna, Eur. Phys. J. B 11, 193 (1999).
- [2] L. Kullmann, J. Kertész, and R.N. Mantegna, Physica A 287, 412 (2000).
- [3] L.E. Reichl, *A Modern Course in Statistical Physics*, 2nd ed. (Wiley, New York, 1998).
- [4] W.-L. Lin, R.F. Engle, and T. Ito, Rev. Finance Stud. 7, 507 (1997).
- [5] N. Vandewalle, Ph. Boveroux, and F. Brisbois, Eur. Phys. J. B 15, 547 (2000).
- [6] C. Schulze, Int. J. Mod. Phys. C 13, 207 (2002).
- [7] A. Lo and A.C. MacKinlay, Rev. Finance. Stud. 3, 175 (1990).
- [8] J. Campbell, A. Lo, and A.C. MacKinlay, The Econometrics of

Financial Markets (Princeton University Press, Princeton, NJ, 1997).

- [9] R. Mantegna and H.E. Stanley, *Econophysics* (Cambridge University Press, Cambridge, England, 2000).
- [10] Y. Liu, P. Gopikrishnan, P. Cizeau, M. Meyer, C. Peng, and H. Eugene Stanley, Phys. Rev. E 60, 1390 (1999).
- [11] E.F. Fama, J. Financ. 25, 383 (1970).
- [12] R. Furth, Ann. Phys. (Leipzig) 53, 177 (1917); G.H. Weiss, Aspects and Applications of the Random Walk (North-Holland, Amsterdam, 1994).
- [13] G. Bonanno, F. Lillo, and R.N. Mantegna, Quantitative Financ.1, 96 (2001).