Analysis of scientific productivity using maximum entropy principle and fluctuation-dissipation theorem

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Using data retrieved from the INSPEC database we have quantitatively discussed a few syndromes of the publish-or-perish phenomenon, including the continuous growth of the rate of scientific productivity, and the continuously decreasing percentage of those scientists who stay in science for a long time. Making use of the maximum entropy principle and fluctuation-dissipation theorem, we have shown that the observed fat-tailed distributions of the total number of papers x authored by scientists may result from the density-of-states function $g(x; \tau)$ underlying the scientific community. Although different generations of scientists are characterized by different productivity patterns, the function $g(x; \tau)$ is inherent to researchers of a given seniority τ , whereas the publish-or-perish phenomenon is caused only by an external field θ influencing researchers.

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I. INTRODUCTION

Nowadays, (...) evaluations of scientists depend on the number of their papers, their positions in lists of authors, and their journals' impact factors. In Japan, Spain, and elsewhere, such assessments have reached formulaic precision. But bureaucrats are not wholly responsible for these changes-we scientists have enthusiastically colluded. What began as someone else's measure has become our (own) goal. (...) [1]. In fact, a number of scientists all over the world feel that research is in crisis. Academics are having to publish or perish. Scientific articles become a valuable commodity both for authors and publishers [2]. The politics of publication does not only concentrate on publishing as valuable articles as possible. Of course, since articles in leading journals certify one's membership in the scientific elite the impact factor of journals matters but also the total number of publications is of great importance since frequent publications allow one to sustain one's career, and are well seen when applying for funds. Authors have to plan when, how, and with whom to publish their results. Quoting Lawrence [1]: The ideal time is when a piece of research is finished and can carry a convincing message, but in reality it is often submitted at the earliest possible moment. (...) Findings are sliced as thin as salami and submitted to different journals to produce more papers. Scientists, who are aware of the publish-or-perish phenomenon warn that research professionalism may be sacrificed in the pursuit of research grants and fame, or simply for fear of loss of a position.

In this paper, using data retrieved from the INSPEC database, we quantitatively analyze two syndromes of the publish-or-perish phenomenon: continuous growth of the rate of scientific productivity and the continuously decreasing percentage of those scientists who stay in science for a long time.

The paper is organized as follows. In the next section we start with a simple examination of scientific productivity distributions for all INSPEC authors together, as it was done by Lotka [3] and Shockley [4]. Then, we study temporal evolution of the scientists. From the whole database we draw *long-life scientists*, i.e., scientists who were doing research for at

least 18 years. Having such a set of scientists we divide it into the so-called *cohorts* including those who started to publish in a given year T (i.e., $T=1975, 1976, \dots, 1987$). We show that unlike the quickly increasing number of all authors listed in the INSPEC database the number of long-life scientists, as characterized by the year of the first publication T, remains almost constant indicating a decreasing percentage of long-life scientists among all researchers. We also show that histograms of scientific productivity N(x;t,T) within T cohorts, measured by the number of articles x, change over time t from almost exponential (when the cohort contains young scientists) to clearly fat-tailed (when the same cohort includes mature researchers). Additionally, we observe that the number of articles produced by a representative of each cohort increases with the square of seniority $\tau = t - T$, i.e., $\langle x \rangle \sim \tau^2$, indicating that each cohort possesses fixed accelera*tion* parameter $a(T) = \frac{\partial^2 \langle x \rangle}{\partial \tau^2}$ which, on its own turn, quickly increases with T. Finally, in Sec. III, we analyze the observed distributions of scientific productivity in terms of equilibrium statistical physics. We show that the fat-tailed histograms N(x;t,T) may result from the inherent density-ofstates function $g(x; \tau)$ characterizing scientific community. We also introduce the parameter $\theta(t,T)$, which has a similar meaning as the inverse temperature β in the canonical ensemble, and describes an external field influencing scientists. The parameter allows us to quantify the effect of publish-orperish phenomenon.

II. SCIENTIFIC PRODUCTIVITY—FUNDAMENTAL RESULTS

In this study we report on the scientific productivity of all authors (over three million) listed in the INSPEC database [5] in the period of 1969–2004. The database, produced by the Institution of Electrical Engineers, provides a few million records indexing scientific articles published world wide in physics, electrical engineering and electronics, computing, and information technology. Although each INSPEC record contains a number of fields (including publication title, clas-



FIG. 1. The figure explains the procedure used in order to retrieve long-life scientists. We assume that an author belongs to the *T* cohort if the period of time that passed between his/her first and last publication fulfills the relation $T_f - T \ge 17$, where T_f is the year of the last publication indexed in our data set. According to the procedure only the first two authors, whose publication history is depicted in the figure, are considered to be long-life *T* scientists.

sification codes, etc.) for our purposes we have retrieved only two of them: authors' names (i.e., names with all initials) and publication year.

Because this method could lead to the misidentification of scientists we have performed the following control analysis. We have selected only those authors from the database, whose names are unique in the sense that there are no other scientists with the same surname. In this way we have reduced the probability that there can be two scientists with the same surname and initials (a chance that two scientists share the same rare surname and the same initials is very low now). Although the sample size reduction was significant (about 80%), the results look similar except of much larger noise. Newman [6] showed that the error introduced by this name misidentification is of the order of a few percent. In the following sections we present results obtained from the whole available data set.

Having the data we were able to discover the initial year of one's scientific activity T (i.e., year of the first publication) and also the cumulative number of his/her publications in the next years. Additionally, from the whole data set we have drawn long-life scientists [i.e., scientists who were productive for at least 18 years (see Fig. 1)], and we have divided them into the so-called T cohorts, with T having the same meaning as previously.

A few important findings on the evolution of the scientific community can be immediately drawn from the simple comparison of the number of all *T* authors and the number of those authors who turned out to be long-life scientists. However, before we discuss how the numbers and their ratio depend on *T*, two limitations of our data should be noted. First, since the INSPEC database does not contain information about articles published before 1969, the initial year of scientific activity *T* for scientists indexed in the database in the early 1970s may be incorrect. That is why, for further analysis we have restricted ourselves to the period starting at *T* = 1975. Second, due to the criterion of 18 years of activity, taken when specifying *T* cohorts, the number of cohorts is limited to 13, respectively, for *T*=1975,1976,...,1987. Keeping in mind the mentioned constraints one can see (Fig.



FIG. 2. Number of all authors listed in the INSPEC database and the number of long-life scientists versus the year of the first publication T.

2) that although the number of all authors listed in the IN-SPEC database increases every year, the number of long-life scientists remains almost constant (the downward trend observed in the 1980s should not be taken into account as it may result from finite-size effects due to the reduction of the period between T+17 and 2004; consider the case of the second author in Fig. 1). The chief conclusion resulting from the above observations is that the percentage of long-life scientists among all scientists monotonically decreases in time (see the inset in Fig. 2).

In the rest of the section we will concentrate on the fundamental features of distributions describing scientific productivity of authors indexed in INSPEC. As a matter of fact, scientific productivity, measured by the number of papers authored, has a long history of study in sociometrics and bibliometrics, with the articles by Lotka [3] and Shockley [4] being famous early examples. Both of these authors found that the number of papers produced by scientists has a *fattailed* distribution, exhibiting both a large number of authors who contributed only a few articles, and a small number of authors who made a very large number of contributions. Being more precise, Lotka (1926) studied a sample of 6891 authors listed in *Chemical Abstracts* during the period of 1907–1916 finding that the number of authors making x publications was described by a power law

$$N(x) \sim x^{-\gamma},\tag{1}$$

with $\gamma \approx 2$, whereas Shockley (1957) investigated the scientific productivity of 88 research staff members at the Brookhaven National Laboratory in the USA finding the lognormal distribution

$$N(x) \sim \frac{1}{s\sqrt{2\pi}x} e^{-(\ln x - m)^2/(2s^2)}.$$
 (2)

In Fig. 3 we have shown on logarithmic scales histograms of the number of papers written by all authors listed in IN-SPEC and all long-life scientists in the database. As expected, both distributions are highly skewed, and their fat tails are due to long-life scientists. One can also see that the distribution of all authors regardless of their seniority is well described by the log-normal distribution (2), which for rea-



FIG. 3. Histograms of the number of papers written by all authors in INSPEC (solid squares) and long-life scientists in the database (open squares). The solid lines represent fits to the data as described in the text: in the left figure, log-normal distribution (gray line) with $m = 0.43 \pm 0.01$ and $s = 1.69 \pm 0.01$, and the distribution composed of two power laws (black lines), one for small and intermediate events ($\gamma = 1.67 \pm 0.01$) and the other for extreme events ($\gamma = 2.87 \pm 0.03$); in the right figure, log-normal distribution (gray line) with $m = 2.68 \pm 0.01$ and $s = 1.27 \pm 0.01$. The inset shows data binned logarithmically.

sons elaborated by Sornette and Cont [7] (see also Refs. [8,9]) may be confused with the distribution having powerlaw tail (1). In Fig. 3, a part of the log-normal fit to our data, we have shown that the distribution composed of two power laws also fits our data very well. Nevertheless, the exponents γ for both regions of the power-law scaling significantly differ from the exponent $\gamma \approx 2$ predicted by Lotka.

The reported studies show that scientists differ enormously in the number of papers they publish. Although, at present the fat-tailed distributions are not so surprising for physicists as they were 20 years ago, the appearance of highly skewed distributions characterizing scientific productivity is still strange since it refers to scientific elite who have undergone a rigorous selection procedure and is expected to be more homogeneous. At the moment, one may, for example, suggest that the noticed differences between scientists may result from the heterogeneity of the analyzed sample (e.g., as is the case in nonextensivity driven by fluctuations [10,11]). To be ahead of these suggestions, in the following we will concentrate on the analysis of T cohorts, as they were characterized at the beginning of this section. Although the approach makes our data more homogeneous, we are aware that it still does not take into account other factors that influence scientific productivity (e.g., access to resources that facilitate research or geopolitical conditions). In the next section we will try to convince the readership that the effect of those omitted factors may be understood in terms of a single function having the same meaning as the density of states in equilibrium statistical physics.

Due to our approach, whatever differences are observed among T scientists can be logically decomposed into only two sorts: (i) life-course differences, which are the effects of biological and social aging, and (ii) cohort differences, which are differences between cohorts at comparable points in career history. According to our knowledge the only similar analysis of scientific productivity was performed by Allison and Stewart [12], who analyzed a sample of U.S. scientists in university departments offering advanced degrees in biology, chemistry, physics, and mathematics. The authors divided the sample into eight age strata by the number of years since Ph.D., representing different cohorts at different points during their career history. Unfortunately, lacking longitudinal



FIG. 4. Histograms of scientific productivity N(x;t,T) characterizing cohorts of long-life scientists, who started to publish in a given year T=1975 or 1985, and $\tau=t-T=6, 12, 18$. (A detailed description of the figure is given in the text.)



FIG. 5. Change of the average productivity $d\langle x \rangle / d\tau$, and the variance $\langle x^2 \rangle - \langle x \rangle^2$ of cohorts' productivity distributions N(x;t,T) versus seniority $\tau = t - T$. Points represent real data retrieved from the INSPEC database, whereas solid lines express numerical fits according to Eqs. (3) and (4). (A detailed description of the figure is given in the text.)

data the authors were only able to observe life-course differences among scientists, assuming that cohort differences are negligible.

In Fig. 4 we have presented how the histogram of scientific productivity N(x;t,T) depends on time t as a T cohort ages. In general, the scenario is the same for all analyzed T cohorts: N(x;t,T) changes from almost exponential (when a cohort contains young scientists) to clearly fat tailed (when the same cohort consists of mature researchers). The results exemplify life-course differences among long-life scientists, and in some sense confirm the so-called hypothesis of *accumulative advantage* [12], which claims that due to a variety of social and other mechanisms productive scientists are likely to be even more productive in the future, whereas those who produce little original work are likely to decline further in their productivity.

In order to examine cohort differences we have analyzed how the average $\langle x \rangle$ and the variance $\langle x^2 \rangle - \langle x \rangle^2$ of the distribution N(x;t,T) depend on the cohort parameter T= 1975,...,1987, and how they change over time t. We have found that the parameters are well-defined increasing functions of time (see Fig. 5)

$$\frac{\partial \langle x \rangle}{\partial \tau} = a\tau + b, \qquad (3)$$

 $\langle x^2 \rangle - \langle x \rangle^2 = A(\tau - B)^C, \tag{4}$

where $\tau = t - T$ and a, b, A, B, C depend on T (see Table I).

At the moment, it is worth mentioning that although our analysis encompasses only 18 initial years of cohorts' history, we have also verified the above relations for 28 years of activity of the oldest 1975 cohort, finding excellent agreement with the results obtained for other cohorts and for the shorter period of time (see the insets in Fig. 5). Nevertheless, one should be aware that even the most productive scientists in his/her declining years slow down their pace of working. According to Zhao [13], the optimal age for scientific productivity is between 25 and 45, reaching the peak for researchers around 37 (i.e., about 18 years since the beginning of their career). Similar findings have also been reported by Kyvik [14], who found that publishing activity reaches a peak in the 45-49-year-old age group and declines by about 30% among researchers over 60 years old. Summing up, in light of previous results on the relation between age and productivity, findings reported in our paper apply to scientists in the most prolific period of their career.

Now, let us briefly comment on the relations (3) and (4). First, note that the linear dependence on seniority τ in Eq. (4) implies that an average representative of each cohort possesses an acceleration parameter *a*, which is fixed during the whole scientific career. Moreover, the parameter increases with *T* (cf. Table I and Fig. 6), certifying that younger (in

and

TABLE I.	Values of parameters	a,b,A,B,C,E,τ_1	for a few T cohorts.	. See Eqs. (3), (4), and (14).

T cohort	а	b	Α	В	С	E	$ au_1$
1975	0.025	0.39	0.06	-1.02	2.86	0.48	-7.24
1977	0.028	0.40	0.03	-1.47	3.09	0.86	-7.49
1979	0.035	0.37	0.06	-0.97	3.00	0.58	-5.38
1981	0.048	0.36	0.01	-2.15	3.50	3.20	-4.60
1983	0.055	0.39	0.01	-2.38	3.63	3.37	-4.53
1985	0.066	0.42	0.04	-1.38	3.26	1.31	-3.64
1987	0.119	0.35	0.07	-1.36	3.25	1.36	-1.80



FIG. 6. Acceleration parameter a and initial velocity b versus cohort parameter T. As previously, points represent data retrieved from INSPEC, whereas solid lines express a trend in the data.

terms of T) scientists are better skilled to produce more papers than their older colleagues at the same point in their scientific career. It is a matter of debate whether the differences in a are due to better adaptation of young people to technological achievements (i.e., computers and the Internet), or they result from the rough competition between researchers, and are one of the syndromes of the publish-orperish phenomenon. In the next section, exploiting relations (3) and (4), we will show that regardless of the reasoning the explanation of accelerated productivity naturally emerges as a result of the treatment of the scientific community by means of methods borrowed from equilibrium statistical physics.

III. THEORETICAL APPROACH TO SCIENTIFIC PRODUCTIVITY—DENSITY-OF-STATES UNDERLYING THE SCIENTIFIC COMMUNITY

In sociometrics, explanations of highly skewed histograms of scientific productivity N(x) (see Fig. 3) are generally of two (not necessarily exclusive) types [15]. The sacred spark (i.e., heterogeneity) hypothesis says that the observed discrepancies in scientific productivity originate in substantial, predominated differences among scientists in their ability and motivation to do creative research, while the accumulative advantage (i.e., reinforcement) hypothesis [12,16] claims that due to a variety of social and other mechanisms, productive scientists are likely to be even more productive in the future. According to the first hypothesis, skewed distributions of hidden attributes characterizing scientists naturally lead to the skewed distribution of productivity, whereas the second hypothesis argues that the observed fat-tailed histogram N(x) results from sophisticated stochastic processes underlying scientific productivity (see, e.g., Refs. [4,17]).

In this section we will present an alternative explanation of the skewed productivity distributions. Since we have already noticed that the fat tail of the distribution P(x)=N(x)/N characterizing the set of all authors listed in IN-SPEC is due to long-life scientists (cf., Fig. 3), in the following we shall only concentrate on distributions P(x;t,T)=N(x;t,T)/N(T) characterizing T cohorts (see Fig. 4). In order to describe the scientific community, we will exploit the maximum entropy principle [18,19], and we will adopt some of the fundamental concepts from equilibrium statistical mechanics (like statistical ensemble, phase space, and density of states). We will also argue that our approach does not contradict the sociological hypothesis mentioned at the beginning of the section.

In physics, the notion of statistical ensemble means a very large number of mental copies of the same system taken all at once, each of which represent a possible state that the real system might be in. When the ensemble is properly chosen it should satisfy the ergodicity condition, which guarantees that the average of a thermodynamic quantity across the members of the ensemble is the same as the time average of the quantity for a single system.

In our approach we will identify a representative of a given T cohort with a physical system, and we will try to describe such a system (i.e., a long-life scientist) in terms of statistical physics. Since (at least now) we do not have access to parallel worlds, in our approach a large group of copies of the same scientist will be replaced with a large set of macroscopically similar long-life scientists, i.e., scientists belonging to the same T cohort, and taken at a given point in their scientific career $\tau = t - T$. Here, the assumption of macroscopic similarity means that the considered scientists are *exposed* to the same external field (influence) $\theta(t,T)$, which forces (motivates) scientists to publish an average number of publications $\langle x \rangle (t, T)$. The external field (influence) θ has the same meaning as the inverse temperature $\beta = (kT)^{-1}$, which determines the average energy $\langle E \rangle$ in the canonical ensemble [20].

Now, suppose that one would like to establish probability distribution $P(\Omega)$ over a given *T* cohort at time *t*, where

$$\Omega = \{y_1, y_2, \dots, y_n\}$$
(5)

stands for states (i.e., microstates) of a single scientist, who belongs to the considered cohort or ensemble. (Let us explain that the parameters y_i are coordinates of a hidden phase space underlying the scientific community, and determining scientific productivity

$$x = x(\Omega) = x(y_1, y_2, \dots, y_n).$$
 (6)

Of course, there exists a number of such parameters, including research field, IQ level, age, number of co-workers, motivation, funds, etc., but as it turns out in the rest of this section a few important findings about our ensembles may be obtained even without detailed knowledge on the parameters.) Due to the maximum entropy school of statistical physics initiated by Edwin T. Jaynes in 1957 [18,19], the best choice for the distribution $P(\Omega)$ is the one that maximizes the Shannon entropy

$$S = -\sum_{\Omega} P(\Omega) \ln P(\Omega), \qquad (7)$$

subject to the constraint

$$\langle x \rangle(t,T) = \sum_{\Omega} P(\Omega) x(\Omega),$$
 (8)

plus the normalization condition

$$\sum_{\Omega} P(\Omega) = 1.$$
 (9)

The Lagrangian for the above problem is given by the expression

$$\begin{aligned} \mathfrak{L} &= -\sum_{\Omega} P(\Omega) \ln P(\Omega) + \alpha(t, T) \bigg(1 - \sum_{\Omega} P(\Omega) \bigg) \\ &+ \theta(t, T) \bigg(\langle x \rangle(t, T) - \sum_{\Omega} x(\Omega) P(\Omega) \bigg), \end{aligned} \tag{10}$$

where the multipliers $\theta(t, T)$ (external field) and $\alpha(t, T)$ are to be determined by Eqs. (8) and (9). Differentiating \mathcal{L} with respect to $P(\Omega)$, and then equating the result to zero one gets the desired probability distribution over the *T* cohort

$$P(\Omega) = \frac{e^{-\theta(t,T)x(\Omega)}}{Z(t,T)},$$
(11)

where Z(t,T) represents the partition function (normalization constant), and

$$Z(t,T) = \sum_{\Omega} e^{-\theta(t,T)x(\Omega)} = e^{\alpha(t,T)+1}.$$
 (12)

Before we proceed further, let us make some comments here about the maximum entropy approach. First, since each T cohort changes over time t a critical reader may bring the validity of our *equilibrium* approach into question. In fact, it is easy to find that the doubts may be reduced to the question if the time evolution of a long-life scientist may be considered as a quasistatic irreversible process. If so (below we discuss the validity of the assumption), the evolution of the considered system (i.e., long-life scientist) corresponds to a sequence of states that are infinitesimally close to equilibrium, and the equilibriumlike approach is well suited to the problem. At the moment let us note that irreversibility does not clash with quasistatisticity.

It is well known that although most of the physical processes are time symmetric or reversible at the microscopic level, reversibility is often not the case at the macroscopic level. The same can be seen in relation to our system (i.e., long-life scientist), where obtaining knowledge or studying (accompanied by the "reversible process" of forgetting) corresponds to microscopic phenomena, whereas publications can be treated as a kind of macroscopic event.

Now, let us go back to the main problem, i.e., if the time evolution of a long-life scientist may be considered as a quasistatic process. In thermodynamics a quasistatic process is a process that happens infinitely slowly. In practice, such processes can be approximated by performing them "very slowly," which means that the characteristic time corresponding to microscopic phenomena underlying the considered process is much shorter than the characteristic time related to macroscopic events. Of course, in the case of longlife scientists the characteristic time characterizing microscopic processes underlying scientific investigation depends on the diligence of scientists, but since publications (macroscopic events) finalize the process of such an investigation one can assume that the required relation between both time scales in fact exists.

At this point we would like to point out that the same maximum entropy approach has been applied in such different applications as urban and population growth, epidemics, or stock market prices [21]. In most of them, despite their apparent nonequilibricity, the assumption of quasistatisticity allows one to derive results that are comparable with real world observations.

On the other hand we have to mention that there exists a maximum entropy formulation of nonequilibrium statistical mechanics [22]. In this formulation one maximizes the path information entropy, which is a measure of our state of ignorance about which microscopic path the system actually follows over time. This approach could be applied in our investigations if we would like to find a path on which our system evolves between two given time frames (t_1 and t_{inf}). Such a path would correspond to a special sequence of papers published by the same author. Because we analyze probability distributions of microstates only in a given point of time (irrespectively of a history and a future of the system) we decided to use the equilibriumlike approach.

The second comment relates to the *ergodicity* of our ensembles. In statistical physics the ergodic hypothesis says that, over long periods of time, the time spent in some region of the phase space corresponding to microstates with the same energy is proportional to the volume of this region, i.e., that all accessible microstates Ω are equally probable over a long period of time. Equivalently, the hypothesis says that time average and average over the statistical ensemble are the same. In the case of long-life scientists, we may only speculate about the underlying phase space, its dimensionality, and coordinates (5). Even if we were able to enumerate most of the significant coordinates characterizing such scientists, surely a part of these coordinates, including, e.g., motivation, would be impossible to quantify. Summarizing, given the above and other difficulties, it appears impossible to verify the ergodic hypothesis for our ensembles, and the question-if ergodicity is fulfilled here-remains open.

Now, having the theoretical framework we are in a position to analyze how the external field $\theta(t,T)$ influencing scientists depends on *T*, and how it changes over time *t*. In order to calculate the parameter we use the fluctuationdissipation relation

$$\langle x^2 \rangle - \langle x \rangle^2 = -\frac{\partial \langle x \rangle}{\partial \theta} = -\frac{\partial \langle x \rangle}{\partial \tau} \left(\frac{\partial \theta}{\partial \tau} \right)^{-1},$$
 (13)

which may be simply derived from $P(\Omega)$ [Eq. (11)]. (Keep in mind that the ensemble averages $\langle x \rangle$ and $\langle x^2 \rangle$, and also θ depend on both *t* and *T*.) At the moment, note that in the previous section we have already found empirical relations corresponding to both sides of the last formula. Inserting the relations (3) and (4) into Eq. (13), after some algebra one obtains



FIG. 7. Main stage: external field (influence) $\theta(t,T)$ versus seniority $\tau = t - T$ for two cohorts T = 1975 and T = 1985. Subset: productivity parameter defined as $\kappa = \theta^{-1}$ versus τ for the same cohorts.

$$\theta(t,T) = -\int_{\tau_0}^{\tau} \frac{a\xi + b}{A(\xi - B)^C} d\xi = E(\tau - B)^{1-C}(\tau - \tau_1) + D,$$
(14)

where parameters a, b, A, B, C, D depend on T, whereas E, τ_1 are functions of these parameters (see Table I).

In Fig. 7 we have presented how the external field $\theta(t,T)$ changes over seniority τ . Since the field conjugates to the cumulative number of publications, its decreasing character indicates that small values of the field correspond to large productivity, and vice versa-large fields induce small productivity. [The inverse of θ , i.e., $\kappa = \theta^{-1}$, stands for a productivity field, which has a more obvious sociological interpretation: larger κ enforces a larger number of papers (see the inset in Fig. 7).] Having in mind the reverse relationship between θ and the number of publications x, one can argue that the constant of integration D in Eq. (14) must be equal to zero. The reasoning behind the statement is the following. Given that the considered long-life scientists never die, still being in the most prolific period of their career, one may simply imagine that in the limit of $\tau \simeq t \rightarrow \infty$ the total number of publications produced by these scientists must approach infinity, which corresponds to $\theta(\infty, T) = 0$, and respectively, D(T)=0.

The above results allow us to further investigate differences between *T* cohorts. Comparing values of the external field $\theta(t,T)$ influencing *T* scientists at the same point $\tau=t$ -*T* in their scientific career, one can show that the field is a decreasing function of *T* (see Fig. 8). (We have also checked that the decreasing character of $\theta(T+\tau,T)$ versus *T* holds for every value of $\tau=1,2,\ldots,18$.) The above stems from the fact that younger (in terms of *T*) scientists publish more than their older colleagues at the same age. The interesting point here is that statistical physics allows one to describe the phenomenon in terms of the changing external field, which leads to accelerated productivity as described in the previous section.

In order to finalize our theoretical approach to scientific productivity we should explain the mutual relationship between the theoretical distribution $P(\Omega)$ [Eq. (11)] and the



FIG. 8. Differences between cohorts. External field $\theta(t,T)$ coupled to the number of publications *x* versus the cohort parameter *T* for $\tau=t-T=9$. The solid line stands for the trend in the empirical data.

empirical distribution P(x;t,T) (see Fig. 4). Thus, since the two distributions apply to the same ensembles there should exist a possibility to cross from one distribution to the other. Such a possibility appears due to the density-of-states function g(x;t,T), which expresses the number of allowed states Ω [cf. Eq. (5)] that scientists may be in, given that the number of publications corresponding to these states equals x [Eq. (6)]. Using the concept of the density of states one can write

$$P[x(\Omega);t,T] = g(x;t,T)P(\Omega), \qquad (15)$$

and, respectively, the empirical function g(x;t,T), correct to the multiplicative factor Z(t,T), may be obtained from the expression

$$\frac{g(x;t,T)}{Z(t,T)} = P(x;t,T)e^{\theta(t,T)x}.$$
(16)

In Fig. 9 we have presented how the empirical density of states g(x;t,T) depends on x. The most striking feature about g(x;t,T) is that it does not depend separately on time t and T, but it depends on their difference $\tau=t-T$ (cf. bunches of curves shown in the figure)

$$g(x;t,T) \equiv g(x;\tau). \tag{17}$$

The above means that the density of states is an inherent characteristic describing researchers of a given seniority τ . It also certifies that the parameter $\theta(t,T)$ [Eq. (14)] has the meaning of an external field, which is only responsible for filling of the corresponding states [Eq. (5)] in the hidden phase space underlying the scientific community. The analogy between our parameter θ and the inverse temperature β in the canonical ensemble is indeed very close. External conditions expressed by the field θ do not change the considered system, which in our case corresponds to a scientist characterized by a given value of τ . They only influence the probability [Eq. (11)] of realization of a state corresponding to a given productivity x [Eq. (6)]. In particular, the findings allow us to say that representatives of younger cohorts usually



FIG. 9. Density-of-states functions $g(x;\tau)$ underlying different T cohorts at different stages of their scientific career τ .

coauthor many more articles than their counterparts (in terms of the same τ) belonging to older cohorts. It means that due to external requirements (which we interpret as publish-orperish phenomenon) representatives of younger cohorts are skilled (forced) to contribute more articles.

Finally, before we proceed to conclusions let us briefly comment on the shape of the function $g(x; \tau)$ (see Fig. 9). The function monotonically decreases for small and quickly increases for large values of x, having the characteristic minimum for intermediate x. One can argue that the corresponding curvature of $g(x; \tau)$ may result from topological requirements imposed by the relation $x(\Omega)$ [Eq. (6)] on the hidden space $\Omega = \{y_1, y_2, \dots, y_n\}$ [Eq. (5)]. A simple but still reasonable example of such a relation is graphically presented in Fig. 10. (Although the figure presents only twoand three-dimensional phase spaces the below reasoning also holds for higher dimensions.) In the figure, the direction of the dashed lines expresses the growing number of publications x, whereas the area of the *n*-dimensional hypersurface is proportional to the number of states $g(x; \tau)$ of a given value of x. As one can see, the hypersurfaces $x(\Omega)$ corresponding to increasing values of x change from convex to concave. The feature leads to the minimum in the density-ofstates function, and has a nice sociological interpretation.

In order to outline the mentioned sociological interpretation, let us assume that all motivators y_i influencing scientific productivity have some minimal values. Such an assumption seems to be natural since one cannot get a salary lower than a certain limit, and it is impossible to possess a negative number of co-workers. On the other hand, there are no upper limits for these parameters. We are not even in a position to guess their units. It follows that for visualization purposes all motivators may be limited to their positive values, as shown in Fig. 10. Now, in order to justify the suggested convex character of the hypersurface $x(\Omega)$ representing small values of x, one can argue that it corresponds to the leading role of one selected motivator y_i , and an insignificant role of other parameters $y_{i\neq i}$. In some sense, such naive thinking on factors influencing scientists is consistent with a common experience stating that in the early stages of a career only one factor creates motivation for scientific activity (e.g., satisfaction). Along with growing x other motivators start to play a role (e.g., recognition and being in power), what may be expressed by the mentioned *convex-to-concave* crossover.

IV. SUMMARY

In this paper we have attempted to provide a quantitative approach to the publish-or-perish phenomenon, which refers to the pressure to constantly publish work in order to further or sustain one's scientific career. Using data retrieved from the INSPEC database we have quantitatively discussed a few syndromes of the phenomenon, including the continuous growth of the rate of scientific productivity, and the continuously decreasing percentage of those scientists who stay in science for a long time. Methods of equilibrium statistical physics have been applied for the analysis. We have shown that the observed fat-tailed distributions of the total number of papers x authored by scientists may result from a specific shape of the density-of-states function $g(x; \tau)$ underlying the scientific community. We have also argued that although different generations of scientists are characterized by different productivity patterns, the function $g(x; \tau)$ is inherent to researchers of a given seniority τ , and the publish-or-perish phenomenon may be quantitatively characterized by the only one-time and generation-dependent parameter θ , which has the meaning of an external field influencing researchers.



FIG. 10. (Color online) Examples of phase trajectories $x(\Omega)$ in the space of scientific motivators $\Omega = \{y_1, y_2, \dots, y_n\}$ resulting in the corresponding shape of $g(x; \tau)$. (A detailed description of the figure is given in the text.)

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- [1] P. A. Lawrence, Nature (London) 422, 259 (2003).
- [2] M. Gad-el-Hak, Phys. Today 57(3), 61 (2004).
- [3] A. J. Lotka, J. Wash. Acad. Sci. 16, 317 (1926).
- [4] W. Shockley, Proc. IRE 45, 279 (1957).
- [5] http://www.iee.org/publish/inspec/
- [6] M. E. J. Newman, Proc. Natl. Acad. Sci. U.S.A. 98, 404 (2001).
- [7] D. Sornette and R. Cont, J. Phys. I 7, 431 (1997).
- [8] U. Frisch and D. Sornette, J. Phys. I 7, 1155 (1997).
- [9] J. Laherrere and D. Sornette, Eur. Phys. J. B 2, 525 (1998).
- [10] G. Wilk and Z. Włodarczyk, Phys. Rev. Lett. 84, 2770 (2000).
- [11] Ch. Beck, Phys. Rev. Lett. 87, 180601 (2001).
- [12] P. D. Allison and J. A. Stewart, Am. Sociol. Rev. 39, 596 (1974).
- [13] B. Jin, L. Li, and R. Rousseau, J. Am. Soc. Inf. Sci. Technol. 55, 544 (2004).

- [14] S. Kyvik, High. Educ. 19, 37 (1990).
- [15] J. R. Cole and S. Cole, Social Stratification in Science (The University of Chicago Press, Chicago, 1973).
- [16] R. K. Merton, Science 159, 56 (1968).
- [17] H. A. Simon, Models of Man, Social and Rational (Hahner, New York, 1957).
- [18] E. T. Jaynes, Phys. Rev. 106, 620 (1957).
- [19] E. T. Jaynes, Phys. Rev. 108, 171 (1957).
- [20] E. T. Jaynes, Where Do We Stand on Maximum Entropy?, in The Maximum Entropy Formalism, edited by R. Levine and M. Tribus (MIT Press, Cambridge, 1979).
- [21] J. N. Kapur, Maximum Entropy Models in Science and Engineering (Wiley, New York, 1989).
- [22] R. C. Dewar, in Non-Equilibrium Thermodynamics and the Production of Entropy: Life, Earth, and Beyond, edited by A. Kleidon and R. Lorenz (Springer Verlag, Berlin, 2006).